# Linear Theory for Filtering Nonlinear Multiscale Systems with Model Error

#### Tyrus Berry

Department of Mathematics The Pennsylvania State University

Collaborator: John Harlim The Pennsylvania State University

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#### Linear Example: Consider a two-dimensional system of SDEs,

$$dx = (a_{11}x + a_{12}y) dt + \sigma_x dW_x,$$
  

$$dy = \frac{1}{\epsilon} (a_{21}x + a_{22}y) dt + \frac{\sigma_y}{\sqrt{\epsilon}} dW_y,$$

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► Given noisy observations z<sub>m</sub> = x(t<sub>m</sub>) + e<sub>m</sub>, e<sub>m</sub> ~ N(0, R) the filtering problem is to estimate the posterior density p(x(t<sub>m</sub>), y(t<sub>m</sub>) | z<sub>1</sub>, z<sub>2</sub>, ..., z<sub>m</sub>).

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- Given the full model above, the Kalman filter gives the optimal posterior estimate, in the sense of minimum variance.
- Now assume the fast variable is unknown and we only have an averaged model  $dX = \alpha X dt + \sigma dW_x$  for the slow variable.
- ► Can we recover p(x(t<sub>m</sub>) | z<sub>1</sub>, z<sub>2</sub>, ..., z<sub>m</sub>) as accurately as the full model?

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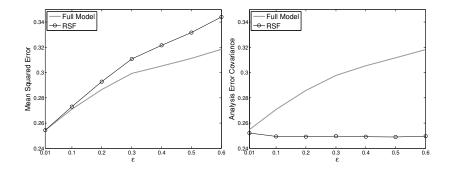
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The **Reduced Stochastic Filter** (RSF) uses the Kalman filter with the averaged model.

## Numerical Results: Model Error with RSF.



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### Understanding covariance inflation

Gottwald & Harlim made the following  $\mathcal{O}(\epsilon)$  closure rigorous.

$$dx = (a_{11}x + a_{12}y) dt + \sigma_x dW_x,$$
  
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Rewrite the fast equation as follows

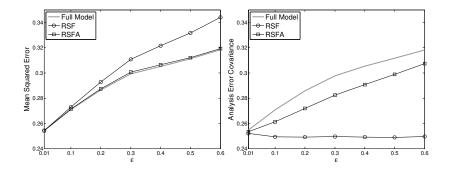
$$y = -\frac{a_{21}}{a_{22}}x - \sqrt{\epsilon}\frac{\sigma_x}{a_{22}}\dot{W}_y + \mathcal{O}(\epsilon)$$

and substitute it to the slow equation and ignore the  $\mathcal{O}(\epsilon)\text{-term},$  we obtain

$$d\tilde{X} = \tilde{a}\tilde{X} dt + \sigma_x dW_x - \sqrt{\epsilon}\sigma_y \frac{a_{12}}{a_{22}} dW_y.$$

**Remarks:** This closure approach is known as the stochastic invariant manifold theory (Fenichel 1979, Boxler 1989).

## Numerical Results: RSF with additive covariance inflation



Improved mean estimates, but the covariance estimates are still underestimated for large  $\epsilon!$ 

The full model steady-state filter covariance  $\hat{S}$  solves,

$$A_{\epsilon}\hat{S} + \hat{S}A_{\epsilon}^{\top} + \hat{S}G^{\top}R^{-1}G\hat{S} + Q_{\epsilon} = 0.$$

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Solving for  $\hat{s}_{11}$  and expanding in  $\epsilon$  we have:

$$-\left(\frac{1+2\epsilon\hat{a}}{R}\right)\hat{s}_{11}^2+2\tilde{a}\left(1+\epsilon\hat{a}\right)\hat{s}_{11}+\left(\sigma_x^2+\epsilon\sigma_y^2\frac{a_{12}^2}{a_{22}^2}\right)+\mathcal{O}(\epsilon^2)=0$$

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The reduced model has steady state covariance solution,  $\tilde{s}$ , that satisfies the 1D Riccati equation,

$$-\frac{\tilde{s}^2}{R}+2\alpha\tilde{s}+\sigma^2=0.$$

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Find parameters  $\{\alpha, \sigma\}$  such that  $\tilde{s} = \hat{s}_{11} + \mathcal{O}(\epsilon^2)!$ 

#### Theorem (Manifold of Parameters, BH2013)

Let  $\hat{s}_{11}$  be the first diagonal component of the 2D algebraic Riccati equation associated with the true filter and let  $\tilde{s}$  be the solution of one-dimensional Ricatti equation associated with the reduced filter. Then  $\lim_{\epsilon \to 0} \frac{\tilde{s} - \hat{s}_{11}}{1} = 0$  if and only if

$$\sigma^2 = 2(\alpha - \tilde{a}(1 - \epsilon \hat{a}))\hat{s}_{11} + \sigma_x^2(1 - 2\epsilon \hat{a}) + \epsilon \sigma_y^2 \frac{a_{12}^2}{a_{22}^2} + \mathcal{O}(\epsilon^2).$$
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**Remarks:** For any parameters on the manifold (1), the reduced filter mean estimate solves,

$$d\tilde{x} = \alpha \tilde{x} dt + \frac{\tilde{s}}{R} (dz - \tilde{x} dt),$$

while the true filter mean estimate for x-variable solves,

$$d\hat{x} = \mathsf{GA}_{\epsilon}(\hat{x},\hat{y})^{ op} dt + rac{\hat{s}_{11}}{R}(dz - \hat{x} dt).$$

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Impose consistency between the actual error covariance,  $\mathbb{E}(e^2)$ , where  $e \equiv \tilde{x} - x$ , and  $\tilde{s}$  to obtain a unique  $\{\alpha, \sigma\}$  in the manifold.

## **Optimal Reduced Stochastic Filter**

## Theorem (Existence and Uniqueness, BH2013) There exists a unique optimal reduced filter given by the following prior model,

$$d\tilde{X} = (\tilde{a} - \epsilon \tilde{a} \hat{a})\tilde{X} dt + \sigma_x (1 - \epsilon \hat{a}) dW_x - \sqrt{\epsilon} \sigma_y \frac{a_{12}}{a_{22}} dW_y, \qquad (2)$$

where  $\tilde{a} = a_{11} - a_{12}a_{21}a_{22}^{-1} < 0$  and  $\hat{a} = a_{12}a_{21}a_{22}^{-2}$ . The optimality is in the sense that, both the mean and covariance estimates converges uniformly to the corresponding estimates from the true filter, with convergence rate on the order of  $\epsilon^2$ .

# **Optimal Reduced Stochastic Filter**

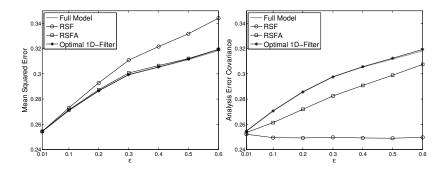
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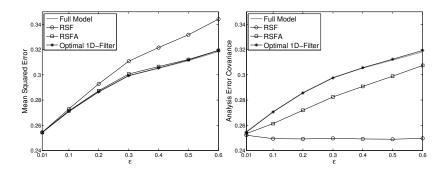
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**Remark:** So, if  $\{\tilde{x}, \tilde{s}\}$  are the solutions of the reduced filter in (2) and  $\{\hat{x}, \hat{s}_{11}\}$  are the solutions of the perfect model, there exists tim-independent constants  $C_1, C_2$ , such that

$$ert \hat{s}_{11}(t) - ilde{s}(t) ert \leq C_1 \epsilon^2,$$
  
 $\mathbb{E}(ert \hat{x}(t) - ilde{x}(t) ert^2) \leq C_2 \epsilon^4.$ 



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#### **Remarks:**

- Notice that for optimal 1D-filter, the MSE (left) approximately equal to the Covariance estimate (right).
   We call a filter **consistent** when the actual error of the mean estimate matches the filtered covariance estimates.
- Optimal solutions are always consistent, but consistent solutions are not necessarily optimal.

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- Finding the reduced model requires imposing consistency on the filter mean and covariance estimates.
- The reduced model includes correction terms in the form of a linear damping and an additive stochastic forcing.
- For general nonlinear filtering problems, it is impractical to find the unique reduced model since it requires imposing consistency on higher-order moments.
- A simple test case shows that general nonlinear problems require multiplicative noise.

#### Nonlinear Filtering Problems

Consider the following prototype continuous-time filtering problem,

$$dx = f_1(x, y; \theta)dt + \sigma_x(x, y; \theta) dW_x,$$
  

$$dy = \frac{1}{\epsilon}f_2(x, y; \theta)dt + \frac{\sigma_y(x, y; \theta)}{\sqrt{\epsilon}} dW_y,$$
  

$$dz = h(x) dt + \sqrt{R}dV.$$

The true filter solutions are characterized by conditional distribution  $p(x, y, t | z_{\tau}, 0 \le \tau \le t)$ , which satisfies the Kushner equation (1964):

$$dp = \mathcal{L}^* p \, dt + p(h - \mathbb{E}[h])^\top R^{-1} (dz - \mathbb{E}[h] \, dt),$$

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#### **Practical issues:**

- ▶ We have no access to *p* for nonlinear problems (SPDE).
- Nonlinearity causes the covariance solutions to depend on higher-order moments and to not equilibrate.

#### Our strategy:

- Do not look for the unique reduced filter since it will require knowledge and consistency of all higher-order moments.
- Pick a simple nonlinear test problem and consider the first two moments of the posterior distribution, p.
- Apply a Gaussian closure to the evolution of these moments.
- Find parameters in a reduced model ansatz by matching the first and second moments of the filtered solutions of perfect model and the reduced models.
- Due to Gaussian closure, even the perfect model may not produce consistent statistics.
- We introduce a consistency metric to determine the performance of the covariance estimate.

#### Definition (Consistency of Covariance)

Let  $\tilde{x}(t)$  and  $\tilde{S}(t)$  be a realization of the solution to a filtering problem for which the true signal of the realization is x(t). The consistency of the realization is defined to be,

$$\mathcal{C}(x, \tilde{x}, \tilde{S}) = \langle \|x - \tilde{x}\|_{\tilde{S}}^2 
angle = rac{1}{n} \langle (x(t) - \tilde{x}(t))^\top \tilde{S}(t)^{-1} (x(t) - \tilde{x}(t)) 
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We say that a filter is consistent if C = 1 almost surely (independent of the realization).

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#### Remarks:

- Consistency does not imply accurate filter.
- A consistency filter with a good estimate of posterior mean has a good estimate of posterior covariance.

Consider [Gershgorin, Harlim, Majda 2010]:

$$\begin{split} \frac{du}{dt} &= -(\tilde{\gamma} + \lambda_u)u + \hat{b} + \tilde{b} + f(t) + \sigma_u \dot{W}_u, \\ \frac{d\tilde{b}}{dt} &= -\frac{\lambda_b}{\epsilon} \tilde{b} + \frac{\sigma_b}{\sqrt{\epsilon}} \dot{W}_b, \\ \frac{d\tilde{\gamma}}{dt} &= -\frac{\lambda_\gamma}{\epsilon} \tilde{\gamma} + \frac{\sigma_\gamma}{\sqrt{\epsilon}} \dot{W}_\gamma, \end{split}$$

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Using the same strategy as for the linear model, we perform asymptotic expansion on the solutions of the optimal filter.

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Using the same strategy as for the linear model, we perform asymptotic expansion on the solutions of the optimal filter.

A detailed computation proves that the optimal reduced filter requires both additive and multiplicative noise.

#### Theorem (Existence, BH13)

Let  $\lambda_u > 0$ , and observations of the full nonlinear test model,  $dz = u dt + \sqrt{R} dV$ . Given identical initial statistics,  $\tilde{u}(0) = \hat{u}(0)$ and  $\tilde{S}(0) = \hat{S}(0) > 0$ , the mean and covariance estimates of a stable continuous-time reduced SPEKF

$$dU = -\alpha U dt + \beta U \circ dW_{\gamma} + \sigma_1 dW_u + \sigma_2 dW_b,$$

with parameters

 $\{\alpha = \lambda_{u}, \beta^{2} = \frac{\epsilon \sigma_{\gamma}^{2}}{\lambda_{\gamma}(\lambda_{u}\epsilon + \lambda_{\gamma})}, \sigma_{1}^{2} = \sigma_{u}^{2}, \sigma_{2}^{2} = \frac{\epsilon \sigma_{b}^{2}}{2\lambda_{b}(\lambda_{b} + \epsilon \lambda_{u})}\} \text{ agree with mean and covariance of a stable continuous-time SPEKF for variable u uniformly, with convergence rate of order-<math>\epsilon$ . Furthermore, the reduced filtered solutions are also consistent, up to order- $\epsilon$ .

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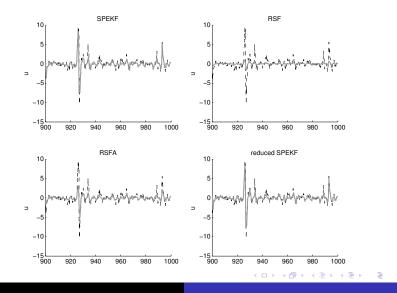
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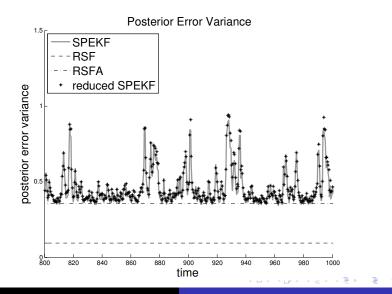
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**Remarks:** Loss an order of  $\epsilon$  accuracy due to multiplicative noise.

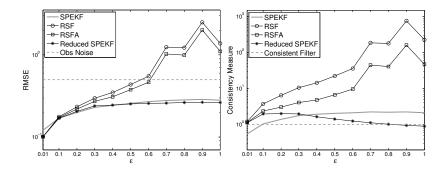
# Numerical solutions in the turbulent transfer energy regime with $\epsilon = 1$ .



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## Numerical Solutions for the nonlinear test filtering problems in a regime that mimics dissipative range



Based on these results, we propose the following ansatz,

$$\left(-\alpha x_i + \sum_{j=1}^N \sigma_{ij} \dot{W}_j + \sum_{j=1}^N \beta_{ij} \circ x_j \dot{V}_j\right)$$

as a stochastic parameterization for model error,

### Example: Strategy for filtering with model errors

Consider the two-layer Lorenz-96 model,

$$\frac{dx_i}{dt} = x_{i-1}(x_{i+1} - x_{i-2}) - ax_i + F + \frac{h_x}{M} \sum_{j=1}^M y_{i,j},$$
  
$$\epsilon \frac{dy_{i,j}}{dt} = y_{i,j+1}(y_{i,j-1} - y_{i,j+2}) - y_{i,j} + h_y x_i,$$

where  $x = x(t) \in \mathbb{R}^N$  and  $y = y(t) \in \mathbb{R}^{NM}$  and the subscript *i* is taken modulo *N* and *j* is taken modulo *M*.

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$$\epsilon \frac{dy_{i,j}}{dt} = y_{i,j+1}(y_{i,j-1} - y_{i,j+2}) - y_{i,j} + h_y x_i,$$

. .

where  $x = x(t) \in \mathbb{R}^N$  and  $y = y(t) \in \mathbb{R}^{NM}$  and the subscript *i* is taken modulo *N* and *j* is taken modulo *M*. **Proposed Reduced Filter Model:** 

$$\frac{dx_i}{dt} = x_{i-1}(x_{i+1} - x_{i-2}) - ax_i + F$$

## Example: Strategy for filtering with model errors

Consider the two-layer Lorenz-96 model,

$$\frac{dx_i}{dt} = x_{i-1}(x_{i+1} - x_{i-2}) - ax_i + F + \frac{h_x}{M} \sum_{j=1}^{M} y_{i,j},$$
  
$$\epsilon \frac{dy_{i,j}}{dt} = y_{i,j+1}(y_{i,j-1} - y_{i,j+2}) - y_{i,j} + h_y x_i,$$

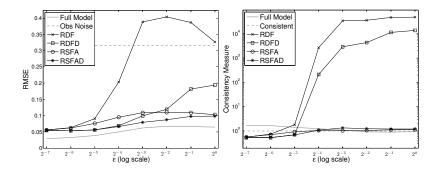
where  $x = x(t) \in \mathbb{R}^N$  and  $y = y(t) \in \mathbb{R}^{NM}$  and the subscript *i* is taken modulo *N* and *j* is taken modulo *M*. **Proposed Reduced Filter Model:** 

$$\frac{dx_i}{dt} = x_{i-1}(x_{i+1} - x_{i-2}) - ax_i + F$$
$$+ \left( -\alpha x_i + \sum_{j=1}^N \sigma_{ij} \dot{W}_j + \sum_{j=1}^N \beta_{ij} \circ x_j \dot{V}_j \right)$$

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- N = 9 slow variables, M = 8 implies 72 fast variables.
- Data generated from the 81-dimensional two-layer L96 model.
- The 9 slow variables are observed with Gaussian noise.
- Ensemble Kalman Filter (EnKF) with each model.
- Parameters  $\alpha$  and  $\sigma$  are fit from the data.
- ▶ We measure the performance of the mean estimate (RMSE).
- We consistency to measure the accuracy of the covariance estimate.
- Consistency  $> 1 \implies$  Underestimating covariance.
- Consistency  $< 1 \implies$  Overestimating covariance.

# Numerical results $(x \in \mathbb{R}^9, y \in \mathbb{R}^{72})$



RDF = Reduced Deterministic Filter ( $\alpha = \beta = \sigma = 0$ ) RDFD = Reduced Deterministic Filter with damping ( $\beta = \sigma = 0$ ) RSFA = Reduced Stochastic Filter with additive noise ( $\alpha = \beta = 0$ ) RSFAD = Reduced Stochastic Filter with damping and additive noise ( $\beta = 0$ )

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